1. Inference in historical sciences
2. State versus Process
3. Importance of Time Scales
4. Maximum Likelihood and Bayesian
5. Example from HIV
   ▶ Getting the tree
   ▶ Using the tree: timing
   ▶ Using the tree: population history
   ▶ Using the tree: counting.
6. Conclusions
Science and repeatability

These steps must be repeatable in order to dependably predict any future results. (Wikipedia)

The Book of Optics by Ibn al-Haytham (1021): conscious reliance upon repeated observations to infer regularities.

Repeated observations can come from:

- Performing controlled experiments, or
- Selecting from a stream of data.

Small correlation structure leads to independent observations.

History may have long correlations: not independent.

- Cosmic Variance: the problem of one universe.
- Large planets have lower density: rule?
- Widespread language similarities: cognitive structure?
- Four limbs: adaptation?
- Similarity across religions: common truth?
Galton’s Problem

Ought we ... to begin by discussing each separate species—man, lion, ox, and the like—taking each kind in hand independently of the rest ... (De partibus Animalium)

In 1889 Sir Edward Tylor presented a paper on correlations between marital systems and societal complexity.

Sir Francis Galton pointed out confounding by borrowing and common descent.


If the cause of the correlation is known, one can reduce it by various methods: data selection, multiple regression, or lagging.

How do we know what is independent?
Synchronic vs. Diachronic

Problem is one of modeling: does synchronic data have enough information about

- Past situations?
- Process rules?

Often, causal structure restricted to a simply connected directed network (tree).

The independence structure of the tree allows us to look at multiple independent realizations.

When tree can be reconstructed, it allows us to deduce diachronic processes from synchronic data alone up to an unknown time scale.

Highly constrained problem: \( n(n - 1)/2 \) distances among \( n \) taxa has only \( 2n - 2 \) independent parameters.
State v Process

‘State’ is
- a statistic of the past;
- sufficient to predict the future.

Changes an integro-differential equation into differential equations:
Future State = Change Rule (Past State)
Change Rule is process.

Time scale separation: State changes faster than Process.
State variables few compared to history:
- Newton’s laws: only position, velocity, and environment.
- State of society: Institutions, norms, knowledge, and myth.
- Linguistics: Language as spoken.

Not always obvious:
- Temperature of floating bodies for motion.
- Remembered poetry on language.
- Family traditions in societal change.
Branching processes have simple causal structure.

Past and present conditionally independent given the state.

State at a branch point (‘node’) splits history into three conditionally independent sectors.

Almost stationary process: the probabilistic rules of change are constant.
Determine the process by fitting to the data.
Timescales

Example:

- Large number of almost independent traits.
- Varying at different rates.
- Rates constant, though different for different traits.
Almost random patterns. No correlations.
Slow Traits

Essentially constant. Replicated initial conditions.
**Informative Traits**

**Intermediate:** Traits that change about once on the timescale
- partition the data
- consistent with a tree
- no correlations except tree concordance

*Excess of traits partitioned consistent with the same tree.*
Hierarchical structuring of non-coextensive traits.
Hierarchical Structure

Fast Random Traits

Hierarchical Structure

Implicational scaling
Stochastic Processes

A stochastic process described by a probability law on states. Usually taken to be infinitely divisible

\[ p(S_0 \to S_t; t) = \sum_{S_{t'}} p(S_0 \to S_{t'}; t') \ p(S_{t'} \to S_t; t - t') \]

Leads to a differential formulation:

\[ \frac{dp(S; t)}{dt} = \sum_{S'} T(S, S'; t) \ p(S'; t) \]

which can be formally solved:

\[ p(S; t) = \sum_{S'} [T \ exp \ \int_{t_0}^{t} T(S, S'; t)dt] \ p(S'; t_0) \]

This allows us to calculate the probability of any set of observations if we are given a model \( T \) and the tree that the data was generated under.
Forward process; Backward inference

Model is constant! How do we determine it?

- Choose model parameters.
- Forward: Either simulate or calculate expected observations.
- Evaluate how expected the actual observations are.
- Backward: Vary model and choose best.

How does this perform with large amount of independent data?

True Model $M$: Independent observations with probability $\{p_i\}$.
Observation $i$ seen with frequency $f_i \approx p_i$.

Let model $M_j$ assign it probabilities $p_{ij}$. Log Total probability

$$\ln \mathcal{L}(M_j) \equiv \ln p(\text{Data} \mid M_j) = \sum_i f_i \ln p_{ij}.$$ 

Maximize over $j$ subject to $\sum_i p_{ij} = 1$: $p_{ij} = f_i$.

The model chosen by Maximum Likelihood is consistent.
Maximum Likelihood Method

Maximum Likelihood Estimate is efficient.

Cramér-Rao inequality: sample-to-sample fluctuation of an estimate $b$ for parameter $\beta$ bounded by:

$$\sigma_b^2 \geq \left( \frac{\partial^2 \ln L}{\partial \beta^2} \right)^{-1} \left\langle \frac{\partial b}{\partial \beta} \right\rangle^2 .$$

Sample fluctuation of an estimate cannot be smaller than product of

- Sensitivity of likelihood to parameter.
- Dependence of estimate on parameter.

Equality often reached by Maximum Likelihood estimate.
Example

Toss a loaded coin many times: how do we determine probability of heads?

As the probability of heads increases, the fraction of heads increases. So, does the number of runs of heads.

Can use different features of the data, e.g.,

- Fraction of heads.
- Average length of runs of heads.

Maximum likelihood chooses fraction of heads, because it is most informative.

Similarly, no need to sort out non-informative fast and slow traits: Maximum Likelihood method weights each trait at its proper time depth.
Model misspecification

Likelihood based methods are ideal
- There is enough data.
- Model in the specified class.

When model does not allow features of the data: one can get bizarre results.
Model: Bengali derived from the Vedic language of India.
We want to find how long back Vedic was spoken.
If we do not recognize that many modern words are actually from Persian, English, Portuguese, Dravidian, Austrasiatic, etc., we will get a very wrong answer.
But, if we allow a probability for random new word: we start getting reasonable results.
Rare forgotten processes can sometimes be replaced by uncorrelated random processes.
Likelihood Ratio

For Normal distribution, \(-2 \ln \mathcal{L} \equiv \chi^2 + \sum \ln(2\pi\sigma^2)\).

- Maximum Likelihood is a generalization of minimum \(\chi^2\).
- Provides confidence intervals.

Adding parameters gives better fit even at random. Best fit models with \(\delta\) more parameters: \(2\Delta \ln \mathcal{L} \sim \chi^2_\delta\).
- Quantifies overfitting.

This can be generalized to Bayesian posterior

\[ p(M_i \mid \text{Data}) \propto p(M_i)\mathcal{L}(M_i) . \]

- Can incorporate prior knowledge.
- Penalizes extra parameters more.
- Can be evaluated by a Markov Chain Monte Carlo.
Example: phylogeny in biology

Can one use these methods to infer history of life? Is history of life tree-like?

But history of what?

Traits are inherited
  ▶ From parents.
  ▶ From peers.
  ▶ From physical environment.
  ▶ From previous changes in environment . . . .

Look for a large co-inherited bundle of traits. Define this as 'vertical transmission.' Other inheritances referred to this baseline.

One such large bundle: genetic traits.
Genotype

Most of life has a strong genotype-phenotype separation.

- Genotype encodes heredity: phenotype is selectable.
- Genotype to phenotype program not easily invertible.
- Genotype changes mainly dictated by chemistry: almost stationary process.

Genotype changes randomly, weakly filtered by selection. Vast amount of almost independent traits, almost stationary process.

- Most of life close to fitness maximum.
- Robustness: Few changes fatal, most neutral.
- High mutation rates harmful. (Eigen’s law: no more than 1 change/unit/generation.)
- Mutation rate maximized for adaptability.
Genetic Change

Most of the changes are ‘point mutations’:

GTAAGACAGTATGATCAGATACTCATAGAAATCTGTGGA → GTAAACAATATGATCAGGTATCTATAGAAATTTTGTTGGA.

Some regions are prone to insertions or deletions.

AGTAATACTACTAGTAAT ↔ ACT ATACTA AAT.

Daughter may form from parts of different parents

...AGGATGGAC... → ...AGGATGCTG...
...TTTATGCTG...
Stationary Independent Sites Model

Consider only point mutations and assume

- **many sites change** at different rates \( r_i \),
- rates are almost constant over time,
- relative probabilities of base substitutions \( T \) the same, and
- sites in the genome change **almost independently**.

\[
\mathcal{L} = \left( \sum_{ik} \overline{p}_i (e^{r_1 x T}) A_i (e^{r_1 y T}) k_i (e^{r_1 v T}) C_k (e^{r_1 w T}) G_k \right) \\
\times \left( \sum_{jl} \overline{p}_j (e^{r_2 x T}) T_j (e^{r_2 y T}) l_j (e^{r_2 v T}) A_l (e^{r_2 w T}) C_l \right),
\]

where \( \overline{p} \) are the initial probabilities.

- Propose reconstructions
- Evaluate reconstructions
- Find history and process that give best reconstructions.
HIV

HIV is a virus about 9719 bases long. More than 1200 almost complete sequences known.

The rates vary from site to site.

Different kinds of changes have different probabilities:
Transversion:Transition CT:Transition GA :: 1:3:4
Model verification

This model correctly predicts the time of the earliest HIV sequence:

![Diagram showing the model verification process](image)

- **A)** Branch Length
- **B)** Test case: the ZR1959 sequence
  - Best fit line, real data excluding ZR59
  - Bootstraps excluding ZR59
  - Data point, all post-1983

- **CONSENSUS** branch
- **1959** point
- **ZR59 branch length**
- **Predicted sampling time**
- **Year** range: 1900 to 2000
Origin of HIV

- Tested two genes:
  - env (gp160): 141 sequences of 2038 bases each.
  - gag: 64 sequences of 1363 bases each.

- Results consistent:
  - env: estimate 1931; 95% CI 1915 – 1941
  - gag: estimate 1934; 95% CI 1869 – 1950

Current results at the earlier end.
Uses of phylogeny: Coalescent Theory

In a population of size $N$, two randomly selected strains had a common parent with probability $1/N$.

Phylogeny provides estimates of relative times of common ancestors.

More ‘coalescenses’ when population size small.

In randomly mixed situations, can estimate populations in the past!
Example: Growth of HIV infections

Study HIV within Democratic Republic of Congo.
Fits the model

\[ N = N_0(\alpha + (1 - \alpha)e^{rt}) \]
Uses of phylogeny: A Problem in HIV immunology

Cytotoxic T Lymphocytes (CTL) recognize and kill virally infected cells.

Small bits of virus presented for recognition by class I Human Leukocyte Antigen (HLA).

Which bits of the virus (epitopes) are recognized by CTLs?
Direct Solution

Find the HLA type of the individual.
Find sequence positions where change is selected over time.
Construct overlapping stretches of small peptides.
Study binding.

Statistical Solution

Individuals differ in HLA.
If an epitope recognized, it may escape by changing.
Study a population and HIV extracted from them.
If an HIV sequence position correlates with host HLA, it is likely to be in an epitope.
Incorrect Results!!!

89/202 (37–51%) positions in the protein Pol correlated with host HLA.

11 (3–10%) significant even after correction for multiple tests.


A separate study: 346/624 (51–59%) positions correlated.

80 (10–16%) significant with a cutoff on false positive rate of 20%.


Conclusion:
CTL escape significant factor in shaping HIV evolution.
Dataset from Perth dominated by B subtype.

All C*1701+ people in the dataset are C subtype infected. C*1701 common in south Africa, south African epidemic dominated by C subtype.

All clades except B are dominated by Leucine.

Valine is not changing to Leucine in C*1701+ people.

Correlation better explained by common descent and migration.
Phylogeny provides a Model of Covariance

In 1985 Felsenstein proposed phylogenetically independent contrasts.

Consider a trait diffusing on a phylogenetic tree:
- The changes on each branch are independent variables, with variance given by branch length.
- The ancestral state can be estimated from the descendants by a mean, weighted by inverse branch lengths.
- A contrast is normalized difference between two daughter nodes:
  
  \[
  \frac{A/\sigma^2_A - B/\sigma^2_B}{1/\sigma^2_A + 1/\sigma^2_B}
  \]
Markovian Processes

In a Markov process, the *changes* at various instants, conditioned by the *state* at that instant, are still independent: so instead of looking at the state, one can look at the change.

This formalizes the intuitive problem noticed before: Valine was not changing to Leucine in C*1701+ people.

So, we devise the following method:

- Calculate the ancestor of sequences
- Select cases with common ancestor.
- Correlate *change* or *not change* with feature.
In other words instead of looking at a table like:

<table>
<thead>
<tr>
<th></th>
<th>V</th>
<th>Not-V</th>
</tr>
</thead>
<tbody>
<tr>
<td>C*1701+</td>
<td>0</td>
<td>7</td>
</tr>
<tr>
<td>Not C*1701+</td>
<td>115</td>
<td>45</td>
</tr>
</tbody>
</table>

\( p = 0.0002 \)

we should look at tables like:

<table>
<thead>
<tr>
<th></th>
<th>V→V</th>
<th>V→Not-V</th>
</tr>
</thead>
<tbody>
<tr>
<td>C*1701+</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>Not C*1701+</td>
<td>75</td>
<td>7</td>
</tr>
</tbody>
</table>

\( p = 1 \)
True correlation

Restricted to the cases when the parent is an Aspargine, the daughter changes when the patient is B*4002+

![Diagram with data points and arrow indicating B*4002]
Sensitivity and Specificity

The new method does correct for the phylogenetic artifact.

Without phylogenetic correction, *silent* mutations correlate at the same rate with host HLA.
Silent: 10/153 (3–12%)
Non-silent: 138/1732 (7–9%)
Some silents are very significant: \( p < 0.00002 \).

With phylogenetic correction 62/80 significant cases due to clade association, and only 7/80 strongly supported.
4/6 were known epitopes, and 2 more were later experimentally found to be true.
Also performs well on synthetic data.
False Positives

Problem in this case was because

- HIV transmitted through social (sexual) contact.
- Human populations are not panmictic.
- HIV transmission clusters correlate with human genetics.

Time scales:

- Humans are only $10^4$ generations old.
- Mutation rate $2.5 \times 10^{-8}$ per base per generation.
- Human genetic decorrelation time: $> 10^8$ generations.
- HIV only $2 \times 10^4$ generations old.
- Mutation rate $2.3 \times 10^{-5}$ per base per generation.
- Most clusters at least a factor of 2 younger.
- HIV decorrelation time $> 10^5$ generations.

Phylogenetics remain important.
Missing true positives

But more generally, it is a question of statistical independence: closely related pieces of evidence overcounted. This can lead to false negatives as well as false positives.

Non-B subtype is predominantly P, not Q, and has no B*4002+ve.

<table>
<thead>
<tr>
<th></th>
<th>not Q</th>
<th>Q</th>
<th>not Q</th>
<th>Q</th>
</tr>
</thead>
<tbody>
<tr>
<td>B*4002+</td>
<td>4</td>
<td>2</td>
<td>4</td>
<td>2</td>
</tr>
<tr>
<td>B*4002-</td>
<td>14</td>
<td>76</td>
<td>4</td>
<td>76</td>
</tr>
</tbody>
</table>

p = 0.01  p = 0.00004
Beyond Biology

Can these methods be applied to other historical sciences?

▶ Is trait inheritance more important than trait genesis?
▶ Are there individuals: stable collections of traits?
▶ Are states easy to define: a closed system?
▶ Is there vertical transmission: coinheritance bundles?
▶ Do the coinherited traits show hierarchical structure?
▶ Are their traits of differing rates?
▶ Are the distances explained by a tree?
▶ Is the change process stationary?

These conditions are probably satisfied in many fields, and underlying laws may be discoverable. Incorrect counting probably common: incorrect deduction of laws. Important to study historical processes quantitatively.